Assessment of the Possibility of Natural Spread of Citrus Psorosis Disease in Texas

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ABSTRACT. Observations in Argentina, Texas and California several years ago on the increase in incidence of citrus psorosis symptoms in the field suggested a possible natural spread of the disease. Specifically, past and more recent observations on the incidence of psorosis in Texas in nucellar Redblush grapefruit trees and in originally virus-free Rio Red grapefruit trees also support the hypothesis that natural transmission occurs. The incidence of psorosis symptoms in four orchards of grapefruit trees was recorded annually over a 3-yr period, rating trees on a scale of 0 (no symptoms), 1 (some bark peeling and/or gumming), 2 (psorosis-like scaling) and 3 (classic psorosis bark-scaling). To assess the possibility of biotic cause of psorosis spread, the spatial arrangement of infected plants was examined at different spatial scales (single tree, groups of plants, and whole plantings) and stochastic simulation models were used to predict the likelihood and contribution of potential exogenous versus endogenous sources of inoculum. Aggregation was observed both at the single tree, and groups scale that decreased through time. Some evidence of relationships among groups of diseased trees was also observed and stochastic models indicated the likelihood of background (exogenous) sources of inoculum. These spatial and spatio-temporal findings are consistent with expected disease spread via vectors although two different possible cases of spread could be identified among the four plots. *Citrus psorosis virus* is the type member of the Ophiovirus genus, and there is evidence that other members are transmitted through the soil by Olpidium brassicae.

The high incidence of citrus psorosis disease in Texas in the 1940s (3) led to the launching of a budwood certification program in 1948 (21). Although many psorosis-free trees were propagated, the program was voluntary and ceased to exist by 1959. However, the use of virus-free budwood sources, and the effects of four tree-killing freezes between 1951 and 1989 are believed to have been responsible for a dramatic reduction in the disease (19). Timmer and Garnsey (22) observed that trees that were of nucellar origin were showing bark scaling symptoms of psorosis, and suggested that a vector may be involved (22). Research in Argentina also provided evidence that psorosis appears to spread naturally in the field (1). More recently, observations again in Texas have suggested natural spread of this disease (18, 20). The characterization of Citrus psorosis virus (CPsV) as the type member of the genus Ophiovirus and demonstration of transmission of another ophiovirus in lettuce by the fungus *Olpidium* brassicae (11), initiated a study on the possibility of transmission of CPsV by a rootinfecting fungus. It should be noted that not all bark-scaling symptoms in citrus are associated with CPsV: this has been found in Spain (16), as well as in Brazil, French Polynesia and elsewhere (17). In our studies, random samples from trees with typical bark scaling symptoms tested positive for psorosis by biological indexing, ELISA and PCR (18). Orchards with infected trees were mapped to examine the spatial patterns of psorosis symptoms and how these patterns changed through time in an attempt to better understand the underlying mechanisms that contribute to the spread of psorosis.

MATERIALS AND METHODS

Data collection and preparation. The incidence of psorosis symptoms in four grapefruit orchards on sour orange (A4, B5, E1 and F1) of the Texas A & M University-Kingsville Citrus Center was recorded annually over a 3-yr period (2002-2004). The number of trees surveyed was 713 (A4), 129 (B5), 539 (E1) and 612 (F1), making a total of 1,993. The disease status of each tree was recorded rating trees on a scale of 0 =no symptoms, 1 = some bark peeling and/or gumming, 2 = small areas psorosis-like scaling, and 3 = typical psorosis bark-scaling and some leaf ringspots. Biological indexing of 13 trees with symptoms selected by nonprobabilistic sampling confirmed psorosis disease was present. ELISA (23 positive/24 trees with leaf ringspot symptoms, and 18/27 with bark scaling) and PCR tests (8 positive/9 trees) of selected trees provided further confirmation. All assays were conducted as described elsewhere (18). There was a bit of uncertainty relative to the lowest disease rating of 1, which could be related to the initial phases of psorosis infection or could be related to other non-viral causes. Therefore, for analysis purposes, two scenarios were considered to partition the ratings into two groups. For scenario 1, a rating of 0 or 1 was considered psorosis-free and a rating of 3 or 4 was considered psorosis infected. For scenario 2, a nondiseased situation was attributed to a rating of 0, and all other ratings of 2, 3, or 4 were considered indicative of psorosis infection. Because bark scaling symptoms can take an average of 12 yr to appear (23), it is probable that some symptomless trees in this study were already infected, but it was assumed that infection would have likely occurred after other trees showed symptoms. A 0 rating is therefore either a healthy tree, or an asymptomatic infected tree.

Spatial analysis. Binary (presence/absence) spatial maps of psorosis were prepared for all assessment dates for each plot. For the first level of spatial hierarchy, ordinary runs analyses were performed on each data set to determine if aggregation existed between adjacent symptomatic trees within rows and across rows with the use of a Visual Basic EXCEL macro (14, 15, and Gottwald, unpublished data). A nonrandom pattern (i.e., aggregation) of symptomatic trees was assumed if the observed number of runs was less than the expected number of runs at P = 0.05.

For the second level of spatial hierarchy, the data were examined for the presence of aggregation at various quadrat sizes. The incidence data for each plot were partitioned into quadrats of four (2 by 2), trees with the use of a Visual Basic EXCEL macro (Gottwald, unpublished data). When data are expressed as disease incidence, the beta-binomial distribution provides the best fit for random conditions (13). Randomness within quadrats was thus assessed via beta-binomial analysis. The beta-binomial index of dispersion D was used to test for the presence of randomness of symptomatic trees at each quadrat size (10, 13). For the beta-binomial index, a large D (>1) combined with a small P(< 0.05)suggests aggregation of symptomatic trees (12).

In the third level of spatial hierarchy, the strength and directionality or orientation of aggregation among quadrats of various sizes containing symptomatic citrus trees for 20 plots were examined with spatial autocorrelation analysis (7). Data were parsed into quadrat of 2×2 , trees. The x, y spatial location and disease incidence of trees within each quadrat size on each assessment date in the individual citrus plots were used as input data. Autocorrelation proximity patterns were calculated consisting of positively (SL+), negatively (SL-), and non-correlated lag positions from which an evaluation of spatial patterns of disease incidence was performed. The size and shape of core and reflected clusters of SL+ were calculated, in which a core cluster is a group of significant, positively correlated (P =0.05), spatial lag distance classes

that form a discrete and contiguous group with the origin (i.e., lag [0.0]) of the autocorrelation proximity pattern; a reflected cluster is a discrete group of two or more contiguous significant positive lag positions discontiguous with the origin and the core cluster. The strength of aggregation is a measure of the saturation of the core clusters with significantly positive lags (i.e., the proportion of lag positions within the extents of the cluster that were significantly positive). Row effects were evaluated as the number of significant lag positions within $_{\mathrm{the}}$ first row (within) or within the first column (across) of the autocorrelation proximity pattern that are contiguous with the origin (2, 8, 23).

Spatio-temporal analysis. Data for the psorosis epidemics were analyzed using the spatio-temporal stochastic model for disease spread which was fitted using Markov-Chain Monte Carlo (MCMC) stochastic integration methods. For a thorough description of the MCMC model, its application and interpretation of results, refer to Gibson (4, 5, 6). The results of the spatio-temporal analysis can be viewed graphically in a two-dimensional parameter space representing a series of 'posterior density' contours of parameter densities. The two parameters represent local (a_{2}) versus background (b) interactions. The parameter bquantifies the rate at which a susceptible individual acquires the disdue to primary infection ease independent of the infected trees in the plot and is therefore is the simple interest or primary infection rate in a spatio-temporal context. For many viruses and other pathogens that are vector transmitted and dispersed, this usually means from sources of inoculum outside of the host population, i.e., plot. However, for soilborne pathogens, it can also represent an increase in disease due to resident inoculum in the soil as well as from other sources such as outside the plot. Whereas a_2 represents the secondary infection rate in a spatio-temporal context, and quantifies the manner in which the infective challenge presented to a susceptible individual by a diseased individual in the population decreases with the distance between them. As a_2 increases, the secondary transmissions occur over shorter ranges and, so long as *b* is not so large that primary infections dominate, disease maps generated by the model exhibit aggregation.

RESULTS AND DISCUSSION

All trees recorded as having symptoms in one year, either had the same rating or higher the following years. Fig. 1 shows the increase in the percentage of trees in each of the four blocks with bark scaling (scenario 2) for the 3-yr survey period.

Spatial arrangement of psorosis-symptomatic trees. The first level of spatial hierarchy examined was the association of symptom status between adjacent psorosis infected trees. Only a small percentage of within-row or across-row tests expressed aggregation of psorosis-infected trees in either direc-



Fig. 1. Percentage of grapefruit trees displaying psorosis bark scaling symptoms in four blocks in south Texas over a 3-yr period.

tion among all of the plot/year combinations of data sets (Table 1). Scenario 2 had a slightly higher proportion of aggregated tests due to the greater disease incidence attributed to this scenario. However, when plots were considered one long line of trees, 41.6 and 58.3% of tests suggested a nonrandom spatial pattern for within- and across-row orientations. Therefore, aggregation was detectable at the individual tree level but was not strong.

The next level of spatial hierarchy to be examined was the association of symptomatic plants within quadrats (groups of 2×2 plants). The interpretation of the values of the beta-binomial index of dispersion (*D*) suggests a spatial structure of symptomatic plants significantly nonrandom for 4 of 9, and 3 of 10 plot/year combinations tested for scenario 1 and scenario 2, respectively, for which the beta-binomial test could be conducted (Table 2). For 33% of the plot/year combinations for both scenarios tested, D values were higher than 1, i.e., aggregated but did not seem to be related to disease incidence. Therefore, aggregation was detectable at the group level also, but like the individual tree level, it was not strong.

The final level of spatial hierarchy examined was the association among groups (quadrats) of trees as estimated by spatial autocorrelation. For the 2×2 quadrat size, small core clusters consisting of only a single adjacent group, were found in seven of 24 tests and core clusters with more than a single adjacent group of psorosis

TABLE 1	L
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ORDINARY RUNS ANALYSIS OF PSOROSIS IN FOUR CITRUS BLOCKS IN SOUTH TEXAS

			Digoago	Ordinary runs				
Plot	Year	Scenario	incidence	Row	Column	Row (all)	Col (all)	
A-4	2002	Scenario 1	0.156338	0/30	1/21	R	Ν	
		Scenario 2	0.5	3/34	1/21	Ν	R	
	2003	Scenario 1	0.232117	1/33	4/21	Ν	Ν	
		Scenario 2	0.550365	1/33	2/21	R	Ν	
	2004	Scenario 1	0.262774	1/33	5/21	Ν	Ν	
		Scenario 2	0.626277	1/33	3/21	R	Ν	
B-5	2002	Scenario 1	0.129496	0/12	1/7	R	R	
		Scenario 2	0.438849	0/19	1/8	Ν	Ν	
	2003	Scenario 1	0.235294	0/17	0/8	R	R	
		Scenario 2	0.602941	0/20	1/8	Ν	Ν	
	2004	Scenario 1	0.235294	0/17	0/8	R	R	
		Scenario 2	0.647059	0/19	2/8	Ν	Ν	
E-1	2002	Scenario 1	0.0909091	0/30	1/9	Ν	Ν	
		Scenario 2	0.307978	4/53	3/10	Ν	Ν	
	2003	Scenario 1	0.103896	0/32	1/10	Ν	Ν	
		Scenario 2	0.376623	4/55	1/10	\mathbf{R}	\mathbf{R}	
	2004	Scenario 1	0.111317	0/33	2/10	Ν	Ν	
		Scenario 2	0.395176	3/55	3/10	R	Ν	
F-1	2002	Scenario 1	0.0359477	0/16	0/10	\mathbf{R}	\mathbf{R}	
		Scenario 2	0.333333	3/48	2/13	R	R	
	2003	Scenario 1	0.0473856	0/19	0/9	\mathbf{R}	\mathbf{R}	
		Scenario 2	0.343137	3/48	2/13	\mathbf{R}	\mathbf{R}	
	2004	Scenario 1	0.0473856	0/19	0/9	R	R	
		Scenario 2	0.370915	3/48	3/13	R	Ν	

Values shown for each plot in each assessment date are the proportion of the number of test rows with significant aggregation (P = 0.05) divided by the total number of rows tested (row with more than 1 diseased tree).

Row (all) and Col (all) tests consider the plot as one long row or column respectively. R = random or non aggregated situation indicated. N = non-random or aggregated situation indicated.

			D:	Beta-binomial parameter (?)	Dispersion Index $(D)^{\rm b}$		
Plot	Year	Disease Tear Scenario incidence p		${\rm Quadrat}\; 2 \times 2^{\scriptscriptstyle \rm a}$	Quadrat 2×2		
A-4	2002	1	0.156338	0.0735507*	1.21186*		
		2	0.5	0.0795065^{*}	1.22791^{**}		
	2003	1	0.232117	0.0488198	1.14063		
		2	0.550365	0.045583	1.14025		
	2004	1	0.262774	0.0201918	1.06869		
		2	0.626277	0.029585	1.10199		
B-5	2002	1	0.129496	0.0957519	1.18322		
		2	0.438849	0.213857	1.46639^{*}		
	2003	1	0.235294	0.12417	1.31036		
		2	0.602941	0.277041^{*}	1.54741^{**}		
	2004	1	0.235294	0.12417	1.31036		
		2	0.647059	0.352268^{*}	1.66835^{***}		
E-1	2002	1	0.0909091	0.131118^{*}	1.3249^{**}		
		2	0.307978	0.0168484	1.06593		
	2003	1	0.103896	0.1047^{*}	1.29181^{*}		
		2	0.376623	0.0138735	1.0532		
	2004	1	0.111317	0.113808^{*}	1.30008^{**}		
		2	0.395176	0.0444073	1.13467		
F-1	2002	1	0.0359477	0	0.880011		
		2	0.333333	0	0.924655		
	2003	1	0.0473856	0	0.839172		
		2	0.343137	0	0.970032		
	2004	1	0.0473856	0	0.839172		
		2	0.370915	0.0548352	1.17827		

TABLE 2 BETA-BINOMIAL PARAMETER AND DISPERSION INDEX FOR PSOROSIS INCIDENCE IN FOUR CITRUS BLOCKS IN SOUTH TEXAS

^aMaximum likelihood estimate of the beta-binomial aggregation parameter θ . Significant departures from zero were determined by a t test, t = θ /s.e.(θ)) and indicated overdispersion. Significance is indicated by *, **, *** at respectively *P* = 0.05, *P* = 0.01 and *P* = 0.001. Values in italic indicate that the likelihood estimation procedure of the p and θ parameters of the beta binomial distribution has failed and that the parameter θ was calculated using the moment method but its departure from zero could not be tested (24).

^bIndex of dispersion (*D*) values for the indicated quadrat size by plot and assessment date for psorosis infected citrus plots in south Texas. Values presented for each assessment date are *D* (=observed variance/binomial variance). Tests for aggregation were performed by comparison of (N-1)x *D* with the chi-square distribution and with the $C(\alpha)$ test (Z statistic) as described in the text. Significance *,**,*** is indicated for the $C(\alpha)$ test. A large (>1) *D* and a small *P* (≤ 0.05) suggest rejection of H_0 (binomial distribution- random pattern of symptomatic trees) in favor of H_1 (overdispersion described by the beta-binomial).

positive trees were found in only two of 24 tests (plot A-4 in 2003 and for E1 in 2004). When core clusters larger that one adjacent lag were indicated, they were incompletely saturated, meaning that they were in loose groups rather than complete compact groups with no missing infected trees (Table 3). There did not seem to be a greater prevalence related to one scenario versus another. Reflected clusters of psorosis-positive trees were found in 9 of 24 plot/year/scenario combinations but were small in size and not prevalent. Where these reflected clusters did exist, they ranged from 13.2 to 51.8 m from the centroid of the main cluster (Table 3).

There was a slightly greater number of aggregations in the within-row direction than the across-row direction and very few edge effects. Thus aggregation was again present but not particularly strong among groups of trees, but there was some evidence of longer distance associations among groups of trees. There also did not appear to

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				Significa	nt lags $^{\rm b}$		c	c		Total	SL+ re	flected clu	sters ^h		Effects ⁱ	
Plot	Year	Scenario	Disease incidence ^a	SL+	SL-	of aggre- gation ^c	cluster size ^d	core cluster saturation ^e	cluster size ^f	reflected clusters ^g	Max. X dist.	Max.Y (dist.	Calc dist. (m)	Within row	Across row	Edge
A-4	2002	1	0.464706	1	0	0	0	NA	1	1	0	6	32.9	1	0	0
		2	0.923529	2	0	0	0	NA	1	2	2	9	36.0	1	0	1S
	2003	1	0.6	7	0	0.714286	5 C	0.555556	2	1	0.5	9	33.1	ũ	1	0
		5	0.923529	က	0	0.333333	1	1	1	7	9	5	51.8	7	0	0
	2004	1	0.676471	5	7	0.2	1	1	4	1	0	3.5	19.2	4	1	0
		5	0.952941	4	0	0.25	1	1	1	က	9	5	51.8	7	1	0
B-5	2002	1	0.333333	0	0	0	0	0	0	0	NA	NA		0	0	0
		2	0.666667	1	0	1	1	1	0	0	NA	NA		1	0	1S
	2003	1	0.458333	0	0	0	0	0	0	0	NA	NA		0	0	0
		2	0.708333	1	0	1	1	1	0	0	NA	NA		0	1	0
	2004	1	0.458333	0	0	0	0	0	0	0	NA	NA		0	0	0
		5	0.708333	1	0	1	1	1	0	0	NA	NA		0	1	0
E-1	2002	1	0.264286	0	0	0	0	0	0	0	NA	NA		0	0	0
		2	0.757143	0	0	0	0	0	0	0	NA	NA		0	0	0
	2003	1	0.307143	1	0	1	1	1	0	0	NA	NA		0	1	0
		5	0.821429	0	0	0	0	0	0	0	NA	NA		0	0	0
	2004	1	0.321429	က	0	1	က	0.6	0	0	NA	NA		1	2	0
		2	0.821429	0	0	0	0	0	0	0	NA	NA		0	0	0
^a Disea:	se incid	ence = Qu ,Y] lags si	adrat diseas gnificantly g	e inciden reater (S	ice. (L+), or]	ess(SL-)this	an expect	ted by chance	e at α = 0.	05 level.						
Core c	luster s	ggregator	i = number of sig	gnificant	SL+ lag	ster / total II is contiguous	with the	e [0,0]]ag pos	sition that	form a dis	crete grou	.p.				
°Core (Reflec:	Jluster ted clus	Saturation ter size =	a = the 'prop the number	ortion' of of contig	SL+ laguous SL	gs contiguou: + in various	s with the clusters	e [0,0]lag pos not contiguo	sition that us with th	form a dis ne core clus	crete grou ter.	ip bounde	d by the ou	itermost l	ag extent:	'n

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"Total number of clusters = the number of contiguous clusters of SL+ in the proximity pattern. "Calculation of distance from centroid of core cluster to centroid of reflected cluster in x and y distance given in tree spaces and Calculated distance in meters. "Effects = the number of SL+ within-row and within-column of the row and column defined by the [0,0] lag. Edge effects are significant (S) or non significant (ns) if

(the number of SL+ at the distal edges of the proximity pattern / total number of SL+) is $\geq 5\%$, and < 5%, respectively.

		ē							s) if
		Edg	0	0	0	0	0	0	ts. ers. int (n
	$Effects^i$	Across row	1	0	2	0	2	0	lag exten ce in meto significa
EXAS		Within row	1	0	1	0	1	0	utermost ed distan (S) or nor
SOUTH T	ısters ^h	Calc dist. (m)	16.4		13.2		13.2		d by the o d Calculat ignificant
OCKS IN	flected clu	Max. Y dist.	2	NA	2	NA	2	NA	ap. 1p. pounde spaces an. ects are s
rRUS BL(SL+ re	Max. X dist.	2	NA	1.5	NA	1.5	NA	crete grou crete grou ter. n in tree s . Edge eff
CTED CI1	Total	reflected clusters ^g	3	0	2	0	7	0	5 level. 6rrm a dis form a dis 5 core clus 2 core clus 2 ance give e [0,0] lag d <5%, res
FOUR PSOROSIS-AFFE	Doffootod	cluster size ^f	1	0	4,2	0	4,2	0	at $\alpha = 0.0$ tion that i tion that i is with the attern. and by thi ined by thi
	1 0000	cluster saturation ^e	NA	0	NA	0	NA	0	ed by chance SL+. SL [0,0]lag posi [0,0]lag posi not contiguou not contiguou a cluster in x 1 column defi d column defi
IS OF FO	0.00 2	cluster size ^d	0	0	0	0	0	0	an expect number of s with the s with the clusters 1 SL+ in th f reflecte to concrete to row and
TION ANALYS	Cturrenth	of aggre- gation ^c	0	0	0	0	0	0	less (SL-) th uster / total n gs contiguou gs contiguou + in various t in various s clusters of t o centroid o to column of th ity pattern /
SLATI(It lags ^{b}	SL-	0	0	0	0	0	0	e. L+), or core clu SL+ lag SL+ lag tous SL tiguou tiguou tiguou tiguou within- virinin proxim
JTOCORREL.	Significar	SL+	3	0	9	0	9	0	 incidence incidence f SL+ in (S grificant grificant of contiguous of contiguou
ATIAL AUT		$\mathbf{Disease}^{-}$ incidence ^a	0.152778	0.833333	0.201389	0.833333	0.201389	0.833333	adrat disease gnificantly g t = number of sig n = the 'propo the number of s = the number s = the number SL+ within- of stal edg
SF		Scenario	1	0	1	2	1	2	and the second
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		Plot	F-1						^a Dise: ^b Num ^b Num ^c Strei ^c Core ^c Core ^c Core ^f Reflei ^f Colci ^b Calci (the n

TABLE 3 (CONTINUED)

be any appreciable difference in the two scenarios which treat the lowest rating of disease as either an indicator of disease or not.

If we consider all three levels of spatial hierarchy examined, we would conclude that psorosis exhibits a weak aggregation within a local area of influence, but that these groups of infected trees seem to be related over distance. Thus, there are weak indications that psorosis is acting as a contagion. Associations among symptomatic trees over longer distances were also indicated to some extent, which could imply the presence of a vector, however an inefficient one, if compared to other vectored pathogens of citrus.

Spatio-temporal stochastic **model.** In the present study we examined the spatio-temporal dynamics of psorosis in four plots using two different scenarios described above. However, unlike the analyses for aggregation at three hierarchical levels where interpretation of the analyses resulted in very similar conclusions concerning aggregation of spatial patterns regardless of scenario, when we examine the spatiotemporal dynamics of the spatial patterns through time, the differences between the two scenarios were quite pronounced. The difference is that for Scenario 1 a very slight amount of bark peeling and gumming was not considered necessarily indicative of psorosis, where these same indicators were considered as indicative of the virus infection for Scenario 2. Thus Scenario 2 considered more plants were infected with psorosis during each assessment date and thus a higher incidence of virus infection. The posterior density contours associated with the four plots for the psorosis pathosystem fell into two general cases and these did not align completely with the two scenarios based on symptoms.

Case 1: Two of the eight posterior density contour maps corresponded to cases where the a_2 was positioned towards the maximum of its range with the highest probability of ≥ 2.5 and *b* was clearly nonzero, usually in the range [0.5, 1.5] but with the highest probability of $\cong 1.0$ (Fig. 2 A, D). A third contour map was related (Fig. 2H), however for this plot/scenario combination, a_2 was also positioned towards the maximum of its range (highest probability estimated to be ≥ 2.5). However, *b* had the highest probability of $\cong 0.5$ but these was also some probability that *b* could also equal 0.

Case 2: In contrast, further inspection of the analogous posterior density maps indicated a second case and revealed a different situation (Fig. 2E). Here the estimated probability resided in a region of parameter space in which *b* is ≥ 0 but less than 1.0 and a_{\circ} assumes values towards the lower end of its range, ≥ 0.5 to 2.1 with a maximum probability of $\cong 1.0$. If we focus on only the highest likelihood region, we see that Figs. 2C and 2B have some similarity but if all including lower probability estimations are taken into account, we see that both a_{2} and b can potentially span the entire range of values and therefore cannot be meaningfully interpreted for both surfaces. Finally for the two posterior density maps not here to for mentioned, the estimated range or probabilities spanned much of the total parameter density space, and are therefore both difficult to interpret and do not provided additional biological interpretation (Fig. 2F, G). Thus we must limit our interpretation to only four of the eight surfaces (Fig. 2A, D, E, H).

Interpretation of these two cases of the spatio-temporal stochastic model results leads to differing conclusions concerning psorosis spread. For case 1, the model suggested that psorosis spread through a combination of random background transmission and a local transmission that operated over short distances and was predominately nearest neighbor and was the predominate situation for three of the four plots studies. If we consider that the background



Fig. 2. Posterior density estimates of Markov-Chain Monte Carlo Simulation of the Spatio-temporal increase of psorosis in citrus plots in south Texas. MCMC posterior density likelihood estimations for local and background influences on disease spread for A,B) Plot A4; C, D) Plot B5; E, F) Plot E1, and Plot F1 (G, H), for Scenarios 1 and 2, respectively. Contour maps represent posterior density estimations L(a) composited over all assessments for the duration of the study for each plot with a minimum of 5% increase in disease incidence between individual assessments.

transmission or primary infection is the result of inoculum sources outside the plots, then this is similar to the spatio-temporal stochastic model results for *Citrus tristeza virus* (CTV) with vectors predominately of the migratory type (9). However, we must also consider that if psorosis is soilborne *via* some mechanism or vector, then we cannot discard the possibility of resident inoculum that is contributing apparently random infections within the plot.

In contrast, for case 2 the results suggested a short-range local infection mechanism which was not restricted to nearest-neighbor interactions. Results also suggested that transmission may have been purely local and that the presence of background infection from sources outside the plot or from soilborne sources was not necessary to explain the observed virus spread. This is similar to the model results for CTV with vectors predominately of the colonizing type (9).

Of course, other interpretations such as a complete lack of a vector are possible. However, if an *Olpidium*-like fungus is associated with the roots of psorosis-infected trees, and the virus can be spread in zoospores, then the two possible cases of virus spread discussed above need to be examined from this perspective. Certainly nearest neighbor and tree to adjacent tree transmission would make sense as virus-infected zoospores could move in irrigation or rain water from tree to tree and even several trees around if some ground flooding occurs. However, transmission across multiple rows especially at oblique angles and over longer distances suggest that other spatial mechanisms in addition to surface water movement would have to come into play. Mechanical spread of zoospore-infected soil and/or zoospore-laden water is one possibility and many chances exist for such an occurrence. The orchards studied were flood irrigated, and banks were made approximately every third row. This means that there is a potential for some soil to be moved each time during irrigation, and that there is some lateral spread of water across rows. We must also consider that at times there was indication of background or primary infection that was not distinguishable for a random resident source of inoculum. During various orchard management practices such as spraying, harvesting, etc., there is potential for mechanical movement of soil on tires, boots, etc. which could deposit isolated new sources of soilborne inoculum. In addition, we also cannot eliminate the possibility of an aerial vector also playing a role as suggested by others (1, 22), nor the possibility that some trees are affected by non-psorosis bark scaling, especially considering that one symptomatic tree did not test positive by ELISA or PCR. Therefore from a spatio-temporal context, for the spread of psorosis in the plots studied the interpretation is unclear, with possible multiple interpretations including aerial and/or soilborne vectors that could contribute to both local and longer distance dispersal of disease.

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